

# Calculating Average Volatility

## 1. Introduction

This paper presents a model for compute average volatility and correlation. Since an arithmetic average of log-normally distributed variables is not log-normal, We approximate the arithmetic average by matching its first and second moments with those of a log-normal variable. *We generate* the volatility of a log-normal variable that approximates an arithmetic average of asset prices. We also calculate the correlation between two log-normal variables chosen to match that between two arithmetic averages of asset prices.

## 2. Definition

We consider an arithmetic average of futures prices, specified with

- a set of  $N$  observation points,  $\{t_i | i = 1, \dots, N\}$ , where  $0 < t_1 < \dots < t_N = T$ ,
- a price at time  $t_i$  of a futures contract maturing at time  $T$ ,  $F_T(t_i)$ ,
- a set of  $N$  asset weights,  $\alpha_i$ , where  $1 \leq i \leq N$ .

The arithmetic average of futures prices is defined by

$$\text{Avg}_1 = \sum_{i=1}^N \alpha_i F_T(t_i) \quad (1)$$

The random variable  $Avg_1$  is not log-normally distributed. However, we approximate the variable  $Avg_1$  with a log-normal variable. Here, the implementation returns the volatility of the approximating log-normal random variable.

Given the same set of  $N$  observation points,  $\{t_i | i = 1, \dots, N\}$ , where  $0 < t_1 < \dots < t_N = T$ , we also consider another arithmetic average of futures prices,  $Avg_2$ , specified with

- price at time  $t_i$  of a futures contract maturing at time  $T$ ,  $V_T(t_i)$ ,
- set of  $N$  asset weights,  $\beta_i$ , where  $1 \leq i \leq N$ .

The function returns the constant instantaneous correlation coefficient between the Brownian motions driving the respective futures prices, to match the correlation between  $Avg_1$  and  $Avg_2$ .

### 3 Valuation Methodology

#### 3.1 The Underlying Process:

We assume that each futures price process  $\{F_T(t) | t \in [0, T]\}$  satisfies a risk-neutral stochastic differential equation of the form

$$\frac{dF_T(t)}{F_T(t)} = \sigma dW_t, \quad (2)$$

where

- $\sigma$  is a constant volatility parameter,
- $W$  is a standard Brownian Motion.

We similarly assume that

$$\frac{dV_T(t)}{V_T(t)} = \varpi dB_t,$$

where

- $\varpi$  is a constant volatility parameter,
- $B$  is a standard Brownian Motion.

Here we assume that  $B$  and  $W$  have a constant, instantaneous correlation coefficient  $\rho$ .

### 3.2 Average Volatility:

We consider the random variable,  $\text{Avg}_1 = \sum_{i=1}^N \alpha_i F_T(t_i)$ , where  $0 \leq t_1 < \dots < t_N < T$ .

$\text{Avg}_1$  is not log-normally distributed. However, we approximate  $\text{Avg}_1$  with a log-normal variable,  $X_T$ , where the process  $X$  satisfies

$$\frac{dX(t)}{X(t)} = \mu_X dt + \sigma_X dW.$$

Our aim is to determine  $\mu_X$  and  $\sigma_X$  by matching the first two moments of  $\text{Avg}_1$  with those of  $X_T$ . This approach (see <https://finpricing.com/FinPricing-ProductBrochure.pdf>) is well known as the Levy approximation. In particular, we consider the system of nonlinear equations

Our aim is to compute  $E[X(T)]$  and  $E[X(T)^2]$ , where  $\left\{ \frac{dX(t)}{X(t)} = \mu_X dt + \sigma_X dW \right.$  **(e1)**

From **(e1)**, we have  $X(T) = X(0) \exp\left(\left(\mu_X - \frac{\sigma_X^2}{2}\right)T + \sigma_X(W_T - W_{t_0})\right)$ ,

$$X(T)^2 = X(0)^2 \exp\left(2\left(\mu_X - \frac{\sigma_X^2}{2}\right)T + 2\sigma_X(W_T - W_{t_0})\right).$$

Using the Laplace Transform property for Gaussian Variables, we get

$$E\left[\exp(\sigma_X(W_T - W_{t_0}))\right] = \exp\left(\frac{\sigma_X^2}{2}(T - t_0)\right).$$

Finally, we have:

- $E[X(T)] = X(t_0) \exp((\mu_X)(T - t_0))$ ,
- $E[X(T)^2] = X(t_0)^2 \exp((2\mu_X + \sigma_X^2)(T - t_0))$ .

Choosing  $X_0 = E[\text{Avg}_1]$ , the solution of the above system is

$$\begin{cases} \mu_X = 0 \\ \sigma_X^2 T = \ln\left(\frac{E[\text{Avg}_1^2]}{(E[\text{Avg}_1])^2}\right), \quad (4) \end{cases}$$

where

- the computation of the first two moments of  $\text{Avg}_1$ ,  $E[\text{Avg}_1]$  and  $E[\text{Avg}_1^2]$ , is completely performed below,
- $\sigma_X$  is the average, or Levy, volatility, which is the value returned by the function.

Let  $F_T(t_i)$  be the price at time  $t_i$  of a futures contract maturing at T.

We consider the random variable,  $\text{Avg} = \sum_{i=1}^N \alpha_i F_{T_i}(t_i)$ ,

where

$F_{T_i}(t_i)$  is the price at time  $t_i$  of a futures contract maturing at  $T_i$ ,

$$0 < t_1 < \dots < t_N,$$

$$t_i < T_i, 1 \leq i \leq N,$$

and

$$F_{T_i}(t_i) = F_{T_i}(t_0) \exp\left(\frac{-(\sigma_{t_i})^2}{2}(t_i - t_0) + \sigma_{t_i}(W_{t_i} - W_{t_0})\right).$$

Our aim is to compute  $E[\text{Avg}]$  and  $E[\text{Avg}^2]$ .

- Using the Laplace Transform property for Gaussian random variables, we directly have:

$$E[\text{Avg}] = \sum_{i=1}^N \alpha_i F_{T_i}(t_0).$$

- On the other hand, we can also express  $\text{Avg}^2$  as

$$\text{Avg}^2 = \sum_{j=1}^N \alpha_j \sum_{i=1}^N \alpha_i F_{T_i}(t_i) F_{T_j}(t_j).$$

Taking the expectation,

$$E[\text{Avg}^2] = \sum_{i=1}^N \sum_{j=1}^N \alpha_i \alpha_j E[F_{T_i}(t_i) F_{T_j}(t_j)],$$

where,

$$E[F_{T_i}(t_i)F_{T_j}(t_j)] = F_{T_i}(t_0)F_{T_j}(t_0) \exp\left(\frac{-\sigma_{t_i}^2}{2}(t_i - t_0) + \frac{-\sigma_{t_j}^2}{2}(t_j - t_0)\right) E\left[\exp(\sigma_{t_i}(W_{t_i} - W_{t_0}))\exp(\sigma_{t_j}(W_{t_j} - W_{t_0}))\right]$$

(e)

Now we set

$$\Delta = E\left[\exp(\sigma_{t_i}(W_{t_i} - W_{t_0}))\exp(\sigma_{t_j}(W_{t_j} - W_{t_0}))\right].$$

We have

$$\Delta = E\left[E\left[\exp(\sigma_{t_i}(W_{t_i} - W_{t_0}))\exp(\sigma_{t_j}(W_{t_j} - W_{t_0}))\right] \middle| \mathfrak{F}_{t_i}\right],$$

where  $\mathfrak{F}_{t_i}$  is the filtration generated by  $\{W_u, 0 \leq u \leq t_i\}$ .

Because  $W_{t_i} - W_{t_0}$  is  $\mathfrak{F}_{t_i}$ -measurable, we have

$$\Delta = E\left[\exp(\sigma_{t_i}(W_{t_i} - W_{t_0}))E\left[\exp(\sigma_{t_j}(W_{t_j} - W_{t_i}))\exp(\sigma_{t_j}(W_{t_i} - W_{t_0}))\right] \middle| \mathfrak{F}_{t_i}\right],$$

then,

$$\Delta = E\left[\exp((\sigma_{t_i} + \sigma_{t_j})(W_{t_i} - W_{t_0}))E\left[\exp(\sigma_{t_j}(W_{t_j} - W_{t_i}))\right] \middle| \mathfrak{F}_{t_i}\right].$$

Now, we use the fact that Brownian Motion has independent increments. Therefore,  $W_{t_j} - W_{t_i}$

is independent of  $\mathfrak{F}_{t_i}$ . This leads us to express  $\Delta$  as:

$$\Delta = E\left[\exp((\sigma_{t_i} + \sigma_{t_j})(W_{t_i} - W_{t_0}))E\left[\exp(\sigma_{t_j}(W_{t_j} - W_{t_i}))\right]\right].$$

We note that the term  $E[\exp(\sigma_{t_j}(W_j - W_{t_i}))]$  is not a random variable, it can be taken out of the bigger expectation.

Therefore, we have

$$\Delta = E[\exp((\sigma_{t_i} + \sigma_{t_j})(W_{t_i} - W_{t_0}))] E[\sigma_{t_j}(W_{t_j} - W_{t_i})].$$

Using again the Laplace Transform property for Gaussian Variables, we finally get:

$$\Delta = \exp\left(\frac{(\sigma_{t_i} + \sigma_{t_j})^2}{2}(t_i - t_0)\right) \exp\left(\frac{\sigma_{t_j}^2}{2}(t_j - t_i)\right).$$

We substitute  $\Delta$  into equation (e) above:

$$E[S_{t_i} S_{t_j}] = F_{T_i}(t_0) F_{T_j}(t_0) \exp(\sigma_{t_i} \sigma_{t_j} (t_i - t_0)), \text{ if } t_i \leq t_j.$$

We can finally conclude:

$$E[\text{Avg}^2] = \sum_{i=1}^N \sum_{j=1}^N \alpha_i \alpha_j E[F_{T_i}(t_i) F_{T_j}(t_j)] = \sum_{i=1}^N \sum_{j=1}^N \alpha_i \alpha_j F_{T_i}(t_0) F_{T_j}(t_0) \exp(\sigma_{t_i} \sigma_{t_j} (\text{Min}(t_i, t_j) - t_0))$$